

In-hospital mortality after hip fracture by treatment setting

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ABSTRACT

Background: Where patients with hip fracture undergo treatment may influence their outcome. We compared the risk of in-hospital death after hip fracture by treatment setting in Canada.

Methods: We examined all discharge abstracts from the Canadian Institute for Health Information with diagnosis codes for hip fracture involving patients 65 years and older who were admitted to hospital with a nonpathological first hip fracture between Jan. 1, 2004, and Dec. 31, 2012, in Canada (excluding Quebec). We compared the risk of in-hospital death, overall and after surgery, between teaching hospitals and community hospitals of various bed capacities, accounting for variation in length of stay.

Results: Compared with the number of deaths per 1000 admissions at teaching hospitals, there were an additional 3 (95% confidence interval [CI] 1–6), 14 (95% CI 10–18) and 43 (95% CI 35–51) deaths per 1000 admissions at large,

medium and small community hospitals, respectively. For the risk of in-hospital death overall, the adjusted odds ratios (ORs) were 1.05 (95% CI 0.99–1.11), 1.16 (95% CI 1.09–1.24) and 1.44 (95% CI 1.31–1.57) at large, medium and small community hospitals, respectively, compared with teaching hospitals. For the risk of post-surgical death in hospital, the adjusted ORs were 1.06 (95% CI 1.00–1.13), 1.13 (95% CI 1.04–1.23) and 1.18 (95% CI 0.87–1.60) at large, medium and small community hospitals, respectively.

Interpretation: Compared with teaching hospitals, the risk of in-hospital death among patients with hip fracture was higher at medium and small community hospitals, and the risk of in-hospital death after surgery was higher at medium community hospitals. No differences were found between teaching and large community hospitals. Future research should examine the role of volume, demand and bed occupancy for observed differences.

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One in 10 patients with hip fracture die during their hospital stay.^{1–3} The risk of death is associated with patient, injury and treatment characteristics.^{4,5} Treatment setting may also influence this risk.^{1,2,6–8} For example, advanced standards of anesthesia and surgery are associated with teaching hospitals,^{9,10} but there is inconsistent evidence for an association between teaching status and in-hospital death.^{9,11–13} Comparing teaching hospitals with community hospitals of different bed capacities may further our understanding of the risk of in-hospital death across treatment settings. Bed capacity is associated with factors of care delivery such as resources, treatment styles and standby capacity.¹⁴

Most patients undergo surgery to repair hip fracture.¹⁵ However, between 6% and 10% of patients do not receive surgery, in some cases because of death while waiting for surgery.^{16,17} To better understand the risk of in-hospital death by treatment setting, outcomes of both surgical and nonsurgical care should be considered. Therefore, we conducted this study to compare

the risks of in-hospital death, overall and after surgery, between teaching hospitals and community hospitals of various bed capacities providing hip fracture care in Canada.

Methods

Design, setting and population

We obtained all discharge abstracts with diagnosis codes for hip fracture (International Classification of Diseases, ninth revision, code 820; and International Statistical Classification of Diseases and Related Health Problems, 10th revision, codes S72.00, S72.01, S72.09, S72.10, S72.19, S72.20) involving patients 65 years and older who were admitted to hospital with a nonpathological first hip fracture between Jan. 1, 2004, and Dec. 31, 2012, in Canada (except for the province of Quebec) from the Canadian Institute for Health Information (CIHI) Discharge Abstract Database.¹⁸ Multiple abstracts with the same patient identifier were combined into a single care episode using the CIHI rules for hospital transfers.^{19,20}

For estimating the risk of postsurgical death, we selected discharge abstracts with procedural codes for hip fracture surgery (Canadian Classification of Health Interventions codes 1VA74^{^^}, 1VA53^{^^}, 1VC74^{^^} and 1SQ53^{^^}; Canadian Classification of Diagnostic, Therapeutic and Surgical Procedures codes 9054, 9114, 9134, 9351, 9359, 9361, 9362, 9363, 9364 and 9369), a valid surgery date and a hospital stay of at least 1 day after surgery. We considered deaths on the day of surgery as intraoperative, and live hospital discharge on the day of surgery as clinically unjustifiable.

Outcomes

The primary outcome was in-hospital death identified by destination code in the discharge abstracts. The time to death was calculated as the number of days from the date of admission (counting the admission day) to the date of death, hospital discharge or 30 days, whichever came first. Postsurgical death referred to deaths on abstracts with a code for hip fracture surgery. The time to postsurgical death was calculated as the number of days from the date of surgery to the date of death, hospital discharge or 30 days, whichever came first. In the analysis of deaths without surgery, we calculated the time to death as the number of days from the date of admission (counting the admission day) to the date of death, surgery, hospital discharge or 30 days, whichever came first.

Treatment setting

We used the definitions of CIHI's Canadian Hospital Reporting Project to classify treatment setting. Members of the Association of Canadian Academic Healthcare Organizations were classified as

teaching hospitals; all other hospitals were community hospitals, grouped by the number of beds: small (< 50 beds), medium (50–199) and large (≥ 200).²¹ Treatment setting at admission was a study variable in the analysis of in-hospital mortality, and treatment setting at surgery was a study variable in the analysis of postsurgical mortality.

Statistical analysis

We used the χ^2 test to compare distributions of patient and care characteristics across treatment settings. We estimated daily rates of death overall and by treatment setting by dividing the number of corresponding events by the total number of inpatient days.

We estimated the cumulative incidence of death as a function of inpatient day, with live discharge as a competing event, assuming patients were at risk of in-hospital death only while they remained in hospital.²² We identified live discharges by the following destination codes: discharged home, discharge to home with support, or transferred to long-term care, palliative care, hospice or addiction treatment. We treated hospital stays that ended by transfer to acute care, discharges on the day after surgery and stays that exceeded 30 days as right-censored observations.²⁰ In the analysis of deaths without surgery, surgery was an additional competing event. We used the Pepe–Mori 2-sample test²² and proportional odds regression models²³ to test whether the cumulative incidences of death differed between teaching hospitals and community hospitals of various bed capacity. The differences were summarized by 30-day risk differences and by odds ratios.²⁴

In the regression analysis, the differences between treatment settings were adjusted for patient age, sex, fracture type, comorbidity (heart failure, chronic obstructive pulmonary disease, acute ischemic heart disease, hypertension, diabetes),^{25,26} province or territory, and the calendar period (2004–2006, 2007–2009 or 2010–2012), day (weekday v. weekend) and time of admission. We adjusted for type (internal fixation v. arthroplasty)²⁷ and timing of surgery in the analysis of postsurgical mortality. We conducted the competing-risk analysis using the pseudo-values method²³ with R packages *cmprsk*,²⁸ *prodlim*²⁹ and *geepack*.³⁰ The number of discharge abstracts was sufficient to detect a 1% increase in the risk of in-hospital death (from 7% to 8%), and in the risk of postsurgical death (from 6% to 7%), with 90% power and a 2-sided significance level of 5%.

Ethics approval

The University of British Columbia Behavioural Research Ethics Board approved this study.

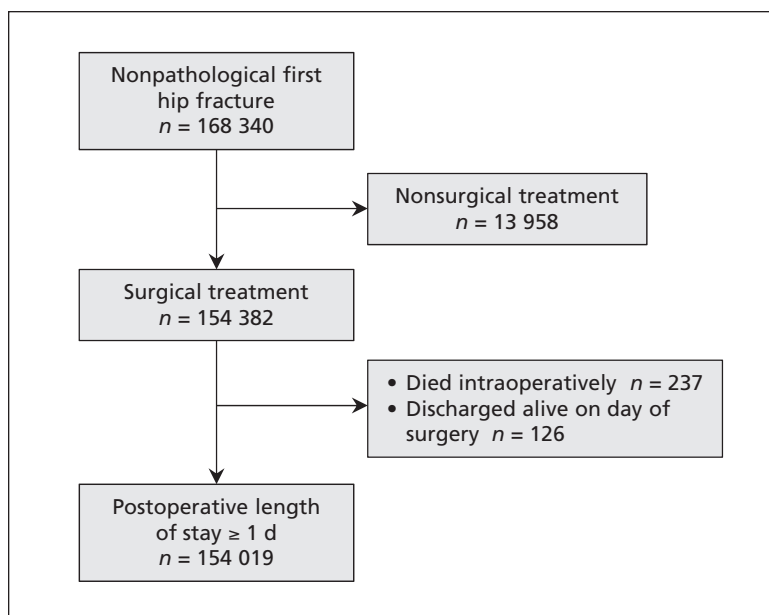


Figure 1: Study population.

Table 1: Patient and care characteristics of 168 340 patients with a first hip fracture, by hospital type at admission

Characteristic	Hospital type; no. (%) of patients				
	All n = 168 340	Teaching n = 58 799	Community, large n = 68 743	Community, medium n = 29 684	Community, small n = 9343
Age at admission, yr					
65–74	25 314 (15.0)	9188 (15.6)	9921 (14.4)	4399 (14.8)	1503 (16.1)
75–84	65 684 (39.0)	22 768 (38.7)	26 991 (39.3)	11 629 (39.2)	3537 (37.9)
85–94	68 154 (40.5)	23 548 (40.0)	28 107 (40.9)	12 093 (40.7)	3763 (40.3)
≥ 95	9188 (5.5)	3295 (5.6)	3724 (5.4)	1563 (5.3)	540 (5.8)
Female sex*	122 696 (72.9)	42 680 (72.6)	50 280 (73.1)	21 746 (73.3)	6684 (71.5)
Fracture type†					
Transcervical	87 248 (51.8)	30 491 (51.9)	35 056 (51.0)	15 644 (52.7)	5180 (55.4)
Pertrochanteric	73 093 (43.4)	25 312 (43.0)	30 360 (44.2)	12 781 (43.1)	3798 (40.7)
Subtrochanteric	7999 (4.8)	2996 (5.1)	3327 (4.8)	1259 (4.2)	365 (3.9)
Calendar year of admission					
2004	18 927 (11.2)	6738 (11.5)	7352 (10.7)	3215 (10.8)	1115 (11.9)
2005	18 971 (11.3)	6703 (11.4)	7424 (10.8)	3285 (11.1)	1122 (12.0)
2006	18 393 (10.9)	6456 (11.0)	7163 (10.4)	3334 (11.2)	1106 (11.8)
2007	18 492 (11.0)	6438 (10.9)	7363 (10.7)	3357 (11.3)	1135 (12.1)
2008	18 446 (11.0)	6427 (10.9)	7580 (11.0)	3263 (11.0)	1079 (11.5)
2009	18 575 (11.0)	6501 (11.1)	7727 (11.2)	3236 (10.9)	1031 (11.0)
2010	18 583 (11.0)	6383 (10.9)	7857 (11.4)	3280 (11.0)	971 (10.4)
2011	18 628 (11.1)	6505 (11.1)	7906 (11.5)	3278 (11.0)	914 (9.8)
2012	19 325 (11.5)	6648 (11.3)	8371 (12.2)	3436 (11.6)	870 (9.3)
Comorbidity‡					
Heart failure	14 250 (8.5)	5048 (8.6)	5709 (8.3)	2488 (8.4)	840 (9.0)
COPD	9315 (5.5)	3117 (5.3)	3604 (5.2)	1816 (6.1)	651 (7.0)
Ischemic heart disease, acute	10 614 (6.3)	3928 (6.7)	4151 (6.0)	1864 (6.3)	550 (5.9)
Cardiac dysrhythmia	16 023 (9.5)	5526 (9.4)	6733 (9.8)	2758 (9.3)	848 (9.1)
Ischemic heart disease, chronic	2955 (1.8)	1048 (1.8)	1240 (1.8)	481 (1.6)	147 (1.6)
Hypertension	9982 (5.9)	3453 (5.9)	4254 (6.2)	1688 (5.7)	495 (5.3)
Diabetes	7862 (4.7)	2877 (4.9)	3125 (4.5)	1325 (4.5)	463 (5.0)
Transferred to another facility	14 273 (8.5)	531 (0.9)	697 (1.0)	6238 (21.0)	6653 (71.2)
Province/territory					
Newfoundland and Labrador	3821 (2.3)	2243 (3.8)	631 (0.9)	783 (2.6)	164 (1.8)
Prince Edward Island	1208 (0.7)	0	1116 (1.6)	33 (0.1)	59 (0.6)
Nova Scotia	7164 (4.3)	2820 (4.8)	1320 (1.9)	2820 (9.5)	204 (2.2)
New Brunswick	5908 (3.5)	1228 (2.1)	2943 (4.3)	1516 (5.1)	189 (2.0)
Ontario	81 556 (48.4)	21 576 (36.7)	39 527 (57.5)	15 563 (52.4)	3163 (33.9)
Manitoba	9625 (5.7)	794 (1.4)	5940 (8.6)	2106 (7.1)	778 (8.3)
Saskatchewan	8802 (5.2)	6686 (11.4)	439 (0.6)	722 (2.4)	955 (10.2)
Alberta	18 330 (10.9)	11 739 (20.0)	4080 (5.9)	634 (2.1)	1876 (20.1)
British Columbia	31 660 (18.8)	11 713 (19.9)	12 747 (18.5)	5404 (18.2)	1792 (19.2)
Northwest, Nunavut or Yukon Territories	266 (0.2)	0	0	103 (0.3)	163 (1.7)
Day of admission					
Weekday	121 857 (72.4)	42 293 (71.9)	49 517 (72.0)	21 722 (73.2)	7052 (75.5)
Weekend	46 483 (27.6)	16 506 (28.1)	19 226 (28.0)	7962 (26.8)	2291 (24.5)
Time of admission§					
0000–0559	23 742 (14.1)	10 844 (18.4)	8829 (12.8)	2989 (10.1)	890 (9.5)
0600–1159	23 632 (14.0)	7046 (12.0)	10 653 (15.5)	4329 (14.6)	1313 (14.1)
1200–1759	56 881 (33.8)	17 809 (30.3)	23 249 (33.8)	11 696 (39.4)	3496 (37.4)
1800–2359	63 987 (38.0)	23 067 (39.2)	25 949 (37.7)	10 669 (35.9)	3643 (39.0)
Type of surgery¶					
Internal fixation	n = 154 382	n = 54 847	n = 64 054	n = 26 576	n = 7302
Arthroplasty	92 445 (59.9)	32 160 (58.6)	38 844 (60.6)	16 006 (60.2)	4411 (60.4)
Arthroplasty	61 937 (40.1)	22 687 (41.4)	25 210 (39.4)	10 570 (39.8)	2892 (39.6)
Timing of surgery¶ **					
Admission day or day after	98 532 (63.8)	32 150 (58.6)	45 476 (71.0)	17 271 (65.0)	2599 (35.6)
≥ 2 d after admission	55 845 (36.2)	22 697 (41.4)	18 576 (29.0)	9305 (35.0)	4703 (64.4)

Note: COPD = chronic obstructive pulmonary disease.

*Does not include 19 patients with unknown sex.

†At admission; for 2084 patients with different fracture types at admission and surgery, the fracture type at surgery is presented.

‡Identified using diagnosis codes from all hospital admissions in the year before the index admission.

§Does not include 98 patients with unknown time of admission.

¶||Only for patients who underwent surgery.

**Does not include 5 patients with unknown timing of surgery.

Results

Patient and care characteristics

A total of 168 340 patients were admitted with a nonpathological first hip fracture between Jan. 1, 2004, and Dec. 31, 2012 (Figure 1). Most (72.9%) were women and almost half (45.9%) were 85 years or older. Fracture type was similarly distributed between transcervical (51.8%) and trochanteric (48.2%) fractures. Overall, 27.9% of the patients had major comorbidities, with cardiac dysrhythmia being the most prevalent (9.5%) (Table 1).

Overall, 58 799 (34.9%) of the patients were admitted to teaching hospitals, and 68 743 (40.8%) were admitted to large, 29 684 (17.6%) to medium and 9343 (5.6%) to small community hospitals (Table 1); type of hospital was unknown for 1771 patients. More patients admitted to small community hospitals (71.2%) were transferred to another facility than were patients admitted to teaching (0.9%), large (1.0%) or medium (21.0%) community hospitals. Admissions between midnight and 0600 were more frequent at teaching hospitals (18.4%) than at large (12.8%), medium (10.1%) or small (9.5%) community hospitals. Weekend admissions were more frequent at teaching hospitals (28.1%) and large community hospitals (28.0%) than at medium (26.8%) or small (24.5%) community hospitals. More patients in Alberta, Saskatchewan, and New-

foundland and Labrador were admitted to teaching hospitals than to large, medium or small community hospitals, compared with patients in other provinces and territories (Table 1).

More patients underwent arthroplasty at teaching hospitals (38.6%) than at large (36.7%), medium (35.6%) or small (31.0%) community hospitals. Of the 154 382 patients who underwent surgery, more underwent surgery on admission day or the day after at large community hospitals (66.2%) than at teaching hospitals (58.6%) or at medium (65.0%) or small (35.6%) community hospitals.

In-hospital mortality

By day 30 after admission, 11 672 (6.9%) hospital stays ended with death, 101 817 (60.5%) ended with live discharge, 26 994 (16.0%) had right-censoring events, and 27 857 (16.6%) stays were longer than 30 days. The average rate of in-hospital death was 4.7 (95% confidence interval [CI] 4.6–4.7) per 1000 patient-days overall, varying from 4.0 (95% CI 3.8–4.1) per 1000 patient-days at teaching hospitals, to 4.8 (95% CI 4.6–4.9), 5.5 (95% CI 5.3–5.8) and 6.3 (95% CI 5.8–6.7) per 1000 patient-days at large, medium and small community hospitals, respectively (Table 2).

Compared with the number of deaths per 1000 admissions at teaching hospitals, there were an additional 3 (95% CI 1–6), 14 (95% CI 10–18) and 43 (95% CI 35–51) deaths per 1000 admissions at large, medium and small community

Table 2: Cumulative incidence of death in hospital and death after surgery, by hospital type

Outcome; hospital type*	No. of patients	No. of deaths†	Rate of death (95% CI)‡	30-d CIF (95% CI)§	Risk difference (95% CI)¶	p value¶¶	Adjusted OR of CIF (95% CI)**
Overall in-hospital mortality††							
Teaching hospital	58 799	3809	4.0 (3.8–4.1)	72 (70–74)	–	–	1.00 (ref)
Community, large	68 743	4739	4.8 (4.6–4.9)	75 (73–77)	3 (1–6)	< 0.01	1.05 (0.99–1.11)
Community, medium	29 684	2199	5.5 (5.3–5.8)	86 (83–90)	14 (10–18)	< 0.001	1.16 (1.09–1.24)
Community, small	9343	799	6.3 (5.8–6.7)	115 (107–123)	43 (35–51)	< 0.001	1.44 (1.31–1.57)
Postoperative mortality‡‡							
Teaching	59 326	2941	3.5 (3.4–3.6)	58 (56–60)	–	–	1.00 (ref)
Community, large	68 921	3694	4.3 (4.2–4.4)	60 (58–62)	4 (0–8)	< 0.01	1.06 (0.99–1.13)
Community, medium	22 986	1234	4.7 (4.5–5.0)	64 (61–68)	11 (5–27)	< 0.001	1.13 (1.04–1.23)
Community, small	1145	70	4.4 (3.4–5.5)	71 (55–87)	10 (4–23)	< 0.05	1.18 (0.87–1.60)

Note: CI = confidence interval, CIF = cumulative incidence function, OR = odds ratio, ref = reference category.
 *Excludes 1771 patients with unknown hospital type at admission.
 †At 30 d from admission for analysis of overall in-hospital mortality; at 30 d after surgery for analysis of postsurgical mortality.
 ‡Per 1000 patient-days.
 §Per 1000 admissions for in-hospital mortality; per 1000 surgeries for postoperative mortality.
 ¶Pepe-Mori test (2-sample test), compared with teaching hospital.
 **Adjusted for age, sex, fracture type, calendar period of admission, comorbidity, province/territory, day of admission, time of admission (also procedure type and time to surgery for postoperative deaths). CIF regression at in-patient days 3, 4, 6, 8, 12, 16, 20, 24 and 30.
 ††Does not include 18 patients with unknown sex.
 ‡‡Does not include 13 958 patients treated nonsurgically, 363 patients discharged on the day of surgery for any reason, and 17 patients with unknown sex or procedure time.

hospitals, respectively (Table 2, Figure 2). For the risk of in-hospital death overall, the adjusted odds ratios (ORs) were 1.05 (95% CI 0.99–1.11), 1.16 (95% CI 1.09–1.24) and 1.44 (95% CI 1.31–1.57) at large, medium and small community hospitals, respectively, compared with teaching hospitals (Table 2).

Postsurgical mortality

For this analysis, we included 154 019 surgically treated patients after excluding patients who died intraoperatively ($n = 237$) or were discharged alive on the day of surgery ($n = 126$) (Figure 1). By day 30 after surgery, 8035 (5.2%) hospital stays ended with death, 95 039 (61.7%) ended with live discharge, 29 324 (19.0%) had right-censoring events, and 21 621 (14.0%) hospital stays were longer than 30 days. The average rate of postsurgical death was 4.0 (95% CI 3.9–4.1) per 1000 patient-days, varying from 3.5 (95% CI 3.4–3.6) at teaching hospitals, to 4.3 (95% CI 4.2–4.4), 4.7 (95% CI 4.5–5.0) 4.4 (95% CI 3.4–5.5) at large, medium and small community hospitals, respectively (Table 2).

Compared with the number of deaths per 1000 surgeries at teaching hospitals, there were an additional 4 (95% CI 0–8), 11 (95% CI 5–27) and 10 (95% CI 4–23) deaths per 1000 surgeries at large, medium and small community hospitals, respectively (Table 2). For the risk of postsurgical in-hospital death, the adjusted ORs were 1.06 (95% CI 0.99–1.13), 1.13 (95% CI 1.04–1.23) and 1.18 (95% CI 0.87–1.60) at large, medium and small community hospitals, respectively, compared with teaching hospitals (Table 2).

Mortality without surgery

For this analysis, we included 13 958 nonsurgically treated patients. By day 30 after admission, 3649 (26.1%) died without surgery, 6778 (48.6%) were discharged without surgery, and 3531 (25.3%) had right-censoring events. The average rate of death without surgery was 6.3 (95% CI 6.1–6.5) per 1000 patient-days, varying from 5.3 (95% CI 5.0–5.6) at teaching hospitals, to 5.9 (95% CI 5.6–6.2), 7.7 (95% CI 7.2–8.3) and 8.6 (95% CI 7.8–9.4) per 1000 patient-days at large, medium and small community hospitals, respectively. Among the 13 958 patients treated nonsurgically, the cumulative incidence of death by inpatient day 30 was 19 (95% CI 18–20) per 1000 admissions at teaching hospitals, and 19 (95% CI 18–20), 29 (95% CI 27–31) and 52 (95% CI 48–57) per 1000 admissions at large, medium and small community hospitals, respectively. Compared with the number of deaths without surgery per 1000 admissions at teaching hospitals, there were an additional 10 (95% CI 8–12) and 34 (95% CI 29–39) deaths per 1000 admissions at medium and small community hospitals,

respectively. There was no difference between teaching and large community hospitals. The adjusted ORs for death without surgery were 1.02 (95% CI 0.92–1.14), 1.50 (95% CI 1.33–1.69) and 2.64 (95% CI 2.30–3.03) at large, medium and small community hospitals, respectively, compared with teaching hospitals.

Interpretation

Compared with teaching hospitals, the risk of in-hospital death was higher at medium and small community hospitals, and the risk of in-hospital death after surgery was higher at medium community hospitals. The difference in postsurgical mortality between teaching hospitals and small community hospitals, although large, was not significant after adjustment. No differences in outcomes were found between teaching hospitals and large community hospitals.

Our findings are consistent with those from previous reports of increased risk of death among patients treated at community hospitals after hip fracture,^{9,11,13} and among patients treated at hospitals with fewer available beds at admission.³¹ As argued elsewhere, the risk of death in hospital also depends on time spent in hospital, which varies by treatment setting.³²

We recently showed a reduction in hospital stay after hip fracture following changes in bed management and changes in policy on access to hip fracture surgery in Canada.²⁰ How these changes were implemented and how effective they were at reducing hospital stay likely varied by treatment setting. Teaching hospitals may shorten stays more

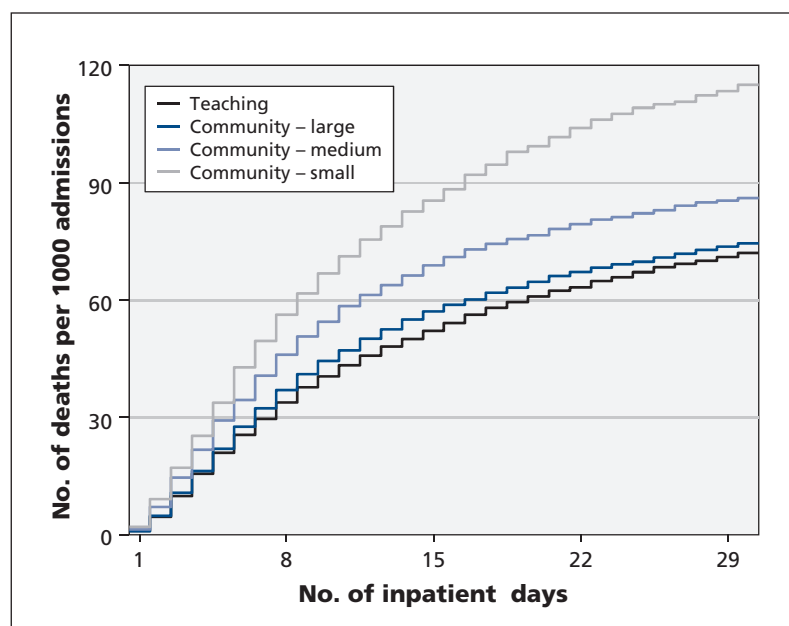


Figure 2: Cumulative incidence of in-hospital death by inpatient days across treatment settings among all patients admitted with first hip fracture.

effectively because discharge options such as rehabilitation and residential care facilities are more prevalent than in community hospitals.³³ Our study accounted for this potential bias. In particular, we used the cumulative incidence to estimate the proportion of patients who died in hospital among all patients admitted to hospital with hip fracture while being exposed to the competing risk of live discharge during the follow-up period.

Postsurgical mortality was higher at medium community hospitals than at teaching hospitals. The difference may be attributable to medium community hospitals having fewer beds, staff and equipment available to ensure access to timely hip fracture care,^{31,34} or to their having a less aggressive treatment style, leaving more patients exposed to potentially fatal immobilized and inflammatory states.^{1,7,35–38} Whether additional resources for medium community hospitals may improve outcomes in this vulnerable patient population requires further investigation.

The risk of death without surgery was higher at medium and small community hospitals than at teaching hospitals. It is not clear whether this difference reflects a need to transfer patients for specialist care not available at medium and small community hospitals. The time required to transfer patients from medium and small community hospitals for care contributes to potentially harmful surgical delay.³⁹ It may be necessary to prioritize these patients on arrival at larger hospitals.

Other structures and processes may influence outcomes of hip fracture care. Previous studies have shown an association between a higher volume of hip fracture surgeries and delays, complications and death.^{40,41} The studies suggest under-prioritization of hip fracture over other surgeries at high-volume sites.^{40,41} Hospital occupancy has also been associated with risk of in-hospital death after hip fracture.³¹ Future research should explore the association among teaching status, bed capacity, occupancy and volume to better our understanding of outcomes of hip fracture care delivery.

Limitations

We conducted a secondary analysis of discharge abstracts with limited variables for adjustment. In particular, patients with hip fracture in different treatment settings may differ by pre-fracture function, level of dependency, injury severity, body composition, cognition, and presence of liver disease, anemia, stroke and secondary hyperparathyroidism.⁴² Further, the abstracts do not provide indication for nonsurgical treatment. Palliative care may have been more frequent at medium and small community hospitals. Classification of treatment settings was based on data from the second

half of the study period.⁴³ This may have led to misclassification of medium and small community hospitals if the number of beds increased across the study years. Bed capacity was not available for teaching hospitals; therefore, we did not investigate difference in mortality by hospital size separately. The hospitals were not identified by their geographic location, which precluded adjustment for urban, rural or remote location. Whether medium and small community hospitals serve more remote populations, or whether Canada's geography could facilitate access to larger hospitals was not factored into our analysis. Few patients underwent surgery at small community hospitals, which, combined with the lack of clinical data, requires some caution in interpretation of the observed differences. Finally, the province of Quebec compiles hospital discharge data in a separate database and does not contribute to the CIHI Discharge Abstracts Database; therefore, the results may not be generalizable to Quebec.

Conclusion

Compared with teaching hospitals, the risk of in-hospital death overall was higher at medium and small community hospitals, and the risk of postsurgical death was higher at medium community hospitals. The difference in postsurgical mortality between teaching hospitals and small community hospitals, although large, was not significant after adjustment. We found no difference between teaching hospitals and large community hospitals. Future research should examine the role of volume, demand and bed occupancy for the observed differences by treatment setting.

References

1. Uzoigwe CE, Burnand HG, Cheesman CL, et al. Early and ultra-early surgery in hip fracture patients improves survival. *Injury* 2013;44:726-9.
2. Neuhaus V, King J, Hageman MG, et al. Charlson comorbidity indices and in-hospital deaths in patients with hip fractures. *Clin Orthop Relat Res* 2013;471:1712-9.
3. Hagino T, Ochiai S, Watanabe Y, et al. Hyponatremia at admission is associated with in-hospital death in patients with hip fracture. *Arch Orthop Trauma Surg* 2013;133:507-11.
4. Penrod JD, Litke A, Hawkes WG, et al. The association of race, gender, and comorbidity with mortality and function after hip fracture. *J Gerontol A Biol Sci Med Sci* 2008;63:867-72.
5. Vestergaard P, Rejnmark L, Mosekilde L. Loss of life years after a hip fracture. *Acta Orthop* 2009;80:525-30.
6. Belmont PJ Jr, Garcia EJ, Romano D, et al. Risk factors for complications and in-hospital mortality following hip fractures: a study using the National Trauma Data Bank. *Arch Orthop Trauma Surg* 2014;134:597-604.
7. Daugaard CL, Jorgensen HL, Riis T, et al. Is mortality after hip fracture associated with surgical delay or admission during weekends and public holidays? A retrospective study of 38,020 patients. *Acta Orthop* 2012;83:609-13.
8. Forte ML, Virmig BA, Swiontkowski MF, et al. Ninety-day mortality after intertrochanteric hip fracture: Does provider volume matter? *J Bone Joint Surg Am* 2010;92:799-806.
9. Weller I, Wai EK, Jaglal S, et al. The effect of hospital type and surgical delay on mortality after surgery for hip fracture. *J Bone Joint Surg Br* 2005;87:361-6.
10. Deiner S, Westlake B, Dutton RP. Patterns of surgical care and complications in elderly adults. *J Am Geriatr Soc* 2014;62:829-35.

11. Taylor DH Jr, Whellan DJ, Sloan FA. Effects of admission to a teaching hospital on the cost and quality of care for Medicare beneficiaries. *N Engl J Med* 1999;340:293-9.
12. Alzahrani K, Gandhi R, Davis A, et al. In-hospital mortality following hip fracture care in southern Ontario. *Can J Surg* 2010; 53:294-8.
13. Yuan Z, Cooper GS, Einstadter D, et al. The association between hospital type and mortality and length of stay: a study of 16.9 million hospitalized Medicare beneficiaries. *Med Care* 2000;38:231-45.
14. Bazzoli GJ, Brewster LR, May JH, et al. The transition from excess capacity to strained capacity in U.S. hospitals. *Milbank Q* 2006;84:273-304.
15. Menzies IB, Mendelson DA, Kates SL, et al. Prevention and clinical management of hip fractures in patients with dementia. *Geriatr Orthop Surg Rehabil* 2010;1:63-72.
16. Technical notes for analyses of hip fracture admissions. Ottawa: Canadian Institute for Health Information; 2005. Available: https://secure.cihi.ca/free_products/WaitTimesReport_tech_Hip_e.pdf (accessed 2016 Oct. 5).
17. Bohm E, Loucks L, Wittmeier K, et al. Reduced time to surgery improves mortality and length of stay following hip fracture: results from an intervention study in a Canadian health authority. *Can J Surg* 2015;58:257-63.
18. Data quality documentation for external users: Discharge Abstract Database, 2010–2011. Ottawa: Canadian Institute for Health Information; 2014.
19. Sheehan KJ, Sobolev B, Guy P, et al. Constructing an episode of care from acute hospitalization records for studying effects of timing of hip fracture surgery. *J Orthop Res* 2016;34:197-204.
20. Sobolev B, Guy P, Sheehan KJ, et al. Time trends in hospital stay after hip fracture in Canada, 2004–2012: database study. *Arch Osteoporos* 2016;11:13.
21. Peer groups in the electronic Discharge Abstract Database reports. Ottawa: Canadian Institute for Health Information; 2015.
22. Pepe MS, Mori M. Kaplan–Meier, marginal or conditional probability curves in summarizing competing risks failure time data? *Stat Med* 1993;12:737-51.
23. Klein JP, Andersen PK. Regression modeling of competing risks data based on pseudovalues of the cumulative incidence function. *Biometrics* 2005;61:223-9.
24. Relative risk, risk differences and odds ratio. Ostend (Belgium): MedCalc. Available: www.medcalc.org/manual/relativerisk_oddsratio.php (accessed 2015 July 24).
25. Nikkel LE, Fox EJ, Black KP, et al. Impact of comorbidities on hospitalization costs following hip fracture. *J Bone Joint Surg Am* 2012;94:9-17.
26. Auais M, Morin S, Nadeau L, et al. Changes in frailty-related characteristics of the hip fracture population and their implications for healthcare services: evidence from Quebec, Canada. *Osteoporos Int* 2013;24:2713-24.
27. Ban I, Palm H, Birkelund L, et al. Implementing, adapting, and validating an evidence-based algorithm for hip fracture surgery. *J Orthop Trauma* 2014;28:e21-6.
28. Gray B. cmprsk: subdistribution analysis of competing risks; 2014. Available: <http://CRAN.R-project.org/package=cmprsk> (accessed 2015 July 24).
29. Gerds T. prodlim: Product-limit estimation for censored event history analysis; 2014. Available: <http://CRAN.R-project.org/package=prodlim> (accessed 2015 July 24).
30. Hojsgaard S, Halekoh U, Yan J. The R package geeppack for generalized estimating equations. *J Stat Softw* 2006;15:1-11.
31. Schilling PL, Campbell DA Jr, Englesbe MJ, et al. A comparison of in-hospital mortality risk conferred by high hospital occupancy, differences in nurse staffing levels, weekend admission, and seasonal influenza. *Med Care* 2010;48:224-32.
32. Nordström P, Gustafson Y, Michaelsson K, et al. Length of hospital stay after hip fracture and short term risk of death after discharge: a total cohort study in Sweden. *BMJ* 2015;350:h696.
33. Williams N, Hardy BM, Tarrant S, et al. Changes in hip fracture incidence, mortality and length of stay over the last decade in an Australian major trauma centre. *Arch Osteoporos* 2013;8:150.
34. Silber JH, Kaestner R, Even-Shoshan O, et al. Aggressive treatment style and surgical outcomes. *Health Serv Res* 2010;45:1872-92.
35. Beloosesky Y, Grinblat J, Pirotsky A, et al. Different C-reactive protein kinetics in post-operative hip-fractured geriatric patients with and without complications. *Gerontology* 2004;50:216-22.
36. Beloosesky Y, Hendel D, Weiss A, et al. Cytokines and C-reactive protein production in hip-fracture-operated elderly patients. *J Gerontol A Biol Sci Med Sci* 2007;62:420-6.
37. Koval KJ, Rust CL, Spratt KF. The effect of hospital setting and teaching status on outcomes after hip fracture. *Am J Orthop* 2011;40:19-28.
38. Peleg K, Savitsky B, Yitzhak B, et al. Different reimbursement influences surviving of hip fracture in elderly patients. *Injury* 2011;42:128-32.
39. Zeltzer J, Mitchell RJ, Toson B, et al. Determinants of time to surgery for patients with hip fracture. *ANZ J Surg* 2014;84:633-8.
40. Kristensen PK, Thillemann TM, Johnsen SP. Is bigger always better? A nationwide study of hip fracture unit volume, 30-day mortality, quality of in-hospital care, and length of hospital stay. *Med Care* 2014;52:1023-9.
41. Metcalfe D, Olufajo OA, Zogg CK, et al. Are older adults with hip fractures disadvantaged in level 1 trauma centers? *Med Care* 2016;54:616-22.
42. Sheehan KJ, Sobolev B, Chudyk A, et al. Patient and system factors of mortality after hip fracture: a scoping review. *BMC Musculoskelet Disord* 2016;17:166.
43. CIHI Portal release note: release 9.14. Ottawa: Canadian Institute for Health Information; 2015.

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