

Risk of fragility fracture among patients with gout and the effect of urate-lowering therapy

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ABSTRACT

BACKGROUND: Previous studies that quantified the risk of fracture among patients with gout and assessed the potential effect of urate-lowering therapy have provided conflicting results. Our study aims to provide better estimates of risk by minimizing the effect of selection bias and confounding on the observed association.

METHODS: We used data from the Clinical Practice Research Datalink, which records primary care consultations of patients from across the United Kingdom. We identified patients with incident gout from 1990 to 2004 and followed them up until 2015. Each patient

with gout was individually matched to 4 controls on age, sex and general practice. We calculated absolute rate of fracture and hazard ratios (HRs) using Cox regression models. Among patients with gout, we assessed the impact of urate-lowering therapy on fracture, and used landmark analysis and propensity score matching to account for immortal time bias and confounding by indication.

RESULTS: We identified 31 781 patients with incident gout matched to 122 961 controls. The absolute rate of fracture was similar in both cases and controls (absolute rate = 53 and 55 per 10 000

person-years, respectively) corresponding to an HR of 0.97 (95% confidence interval 0.92–1.02). Our finding remained unchanged when we stratified our analysis by age and sex. We did not observe statistically significant differences in the risk of fracture among those prescribed urate-lowering therapy within 1 and 3 years after gout diagnosis.

INTERPRETATION: Overall, gout was not associated with an increased risk of fracture. Urate-lowering drugs prescribed early during the course of disease had neither adverse nor beneficial effect on the long-term risk of fracture.

Gout is the most common type of inflammatory arthritis, affecting 2.4% of adults in the United Kingdom.¹ It has been hypothesized that, in common with other chronic inflammatory arthritides such as rheumatoid arthritis and spondyloarthritis,^{2,3} gout may be associated with an increased risk of fracture, primarily owing to the negative effects of chronic inflammation on bone, because proinflammatory cytokines are known to induce bone loss. However, the effects of serum urate on bone health are still under debate^{4,5} and previous studies that have assessed the impact of gout and urate-lowering therapy on fracture risk have provided conflicting results. For instance, a population-based study from Taiwan⁶ found a 17% increased risk of fracture among patients with gout and reported lower fracture risk among those prescribed urate-lowering therapy. In contrast, a registry-based study from Denmark⁷ found a 9% higher risk of fracture among patients prescribed allopurinol compared with patients who were not. A study in the United States⁸ concluded that gout has no effect on the risk of non-vertebral fracture. Two of the studies were based on large administrative data^{6,7} but failed to take into account lifestyle-related factors such as body mass index (BMI) and alcohol consumption; they also

did not adequately address selection bias associated with potentially delayed urate-lowering therapy after diagnosis,⁹ which may have affected their overall conclusions. Fragility fractures are associated with increased cost of health care^{10,11} and a substantial cause of morbidity. Therefore, the aim of this study was to quantify precisely the risk of fracture among patients with gout in the UK and assess the potential effect of urate-lowering therapy on fracture risk estimates using a large population-based primary health care database.

Methods

Data source, design and setting

We used the Clinical Practice Research Datalink,¹² a large database containing UK primary care medical records of anonymized patients (supplementary data, section on Clinical Practice Research Datalink [Appendix 1, available at www.cmaj.ca/lookup/suppl/doi:10.1503/cmaj.170806/-/DC1]). The database is representative of the general UK population in terms of age, sex, ethnicity and lifestyle-related characteristics.^{13,14} We identified individuals with a first-ever recorded Read code diagnosis of gout from general practices between 1990

and 2004, who were then followed up until 2015. The diagnosis of gout was based on a medical code assigned by the physician, which has been previously validated in the Clinical Practice Research Data-link and has a positive predictive value of 90%.¹⁵ Each patient was assigned an index date that corresponded to the date of their diagnosis of gout and randomly matched to 4 controls, who did not have a diagnosis of gout or evidence of urate-lowering therapy, who were registered at the same practice and were alive and contributing data at the index date. They were matched on age (± 3 yr) and sex. Controls were assigned the same index date as their matched gout case. For both cases and controls, follow-up commenced from the index date. Those with a history of fragility fracture, less than 1 year of follow-up before the index date or less than 3 years of follow-up after index date were excluded from the study.

The event of interest was time from the index date until the first diagnosis of fracture. Medical codes for fractures at sites of major osteoporotic fracture were selected (vertebrae, humerus, wrist and hip) in addition to codes for fragility fractures of unspecified site. For the purpose of this study, we were interested only in the incidence of first fracture; thus, all subsequent fracture events were ignored. Van Staa and colleagues¹⁶ carried out external validation of fracture diagnosis in the Clinical Practice Research Data-link and found that 88% and 91% of vertebral and hip fracture diagnoses, respectively, were verified by physicians.

For each individual in our study, we extracted information on relevant lifestyle-related characteristics (smoking status and alcohol consumption), BMI and comorbidities (defined using the Charlson index).¹⁷ We also extracted information on selected medication use (antihypertensive agents, antidiabetic agents, opioids, glucocorticoids, proton pump inhibitors, selective serotonin reuptake inhibitors and bisphosphonates) and history of falls. We ascertained information on comorbidities and lifestyle-related characteristics within 5 years, and on medication use within 1 year before the index date. As the timing of beginning urate-lowering therapy varies after diagnosis of gout, we used landmark analysis to examine the effect of urate-lowering therapy on the risk of first fracture among patients with gout.¹⁸ This method deals with immortal time bias, which biases the results in favour of the treatment under study, by granting a false survival advantage to the treated group. In landmark analysis,¹⁹ a fixed time after the start of therapy is selected for conducting survival analysis (Figure 1). Only patients who were alive and contributing

data at landmark time were included in the analysis. The exposure (urate-lowering therapy) was evaluated between the index date and the landmark time, whereas a fracture event was considered only after the landmark time point. Two landmark points were considered in the analysis (1 and 3 yr after diagnosis), based on a previously published study.²⁰ Only patients who were prescribed more than 6 months of urate-lowering therapy were considered to be exposed.

Statistical analysis

We calculated the incidence of fracture as the number of first recorded fractures per 10000 person-years. Using a Cox regression model, we calculated hazard ratios (HRs) and 95% confidence intervals (CIs), comparing the risk of fracture between gout cases and controls adjusted for various covariates. We accounted for clustering by practice by using robust standard errors. We imputed missing values of BMI by multiple imputation using chained equations. The proportional hazard assumption was tested using Schoenfeld residuals. We stratified our analysis by age, sex and fracture site. To assess the effect of urate-lowering therapy on fracture risk among those with gout, we used propensity score matching methods to account for confounding by indication. The propensity score for urate-lowering therapy represents the probability that a patient is prescribed the therapy for at least 6 months during the exposure window, given their observed covariates (described further in Appendix 1). A logistic regression model was used to estimate propensity scores; subsequently, each patient with gout who was exposed to urate-lowering therapy was matched to 1 patient with gout who had not been exposed to the therapy, based on their propensity score with caliper distance of 0.2.¹⁹ We used a greedy algorithm to select matches: i.e., we selected the closest matching first, then the closest remaining matching, until there were no acceptable matches. We separately compared the risk of fracture among those who received more than 6 months of urate-lowering therapy within a 1- and 3-year exposure window compared with their matched controls who had not been exposed to urate-lowering therapy during that period, using a Cox regression model. We conducted all analyses using Stata version 14.

Ethics approval

This study was approved by the Clinical Practice Research Data-link Independent Scientific Advisory Committee (reference number 15 165RA).

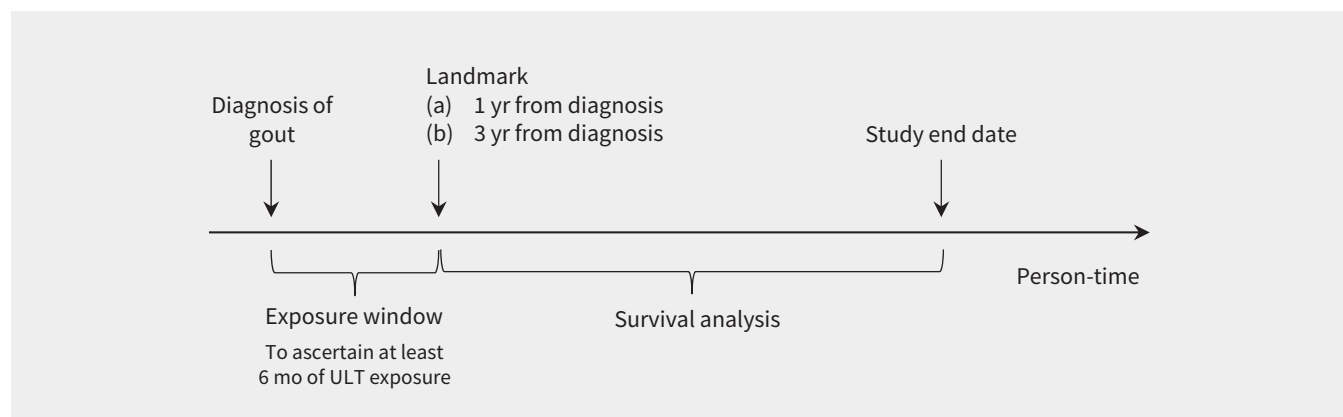


Figure 1: Graphical representation of landmark analysis. Note: ULT = urate-lowering therapy.

Results

We identified 31 781 cases of patients with incident gout who were matched to 122 961 controls. The median follow-up for our study was 10.8 years (interquartile range [IQR] 6.8–13.6 yr). The characteristics of the study population are summarized in

Table 1. A total of 8934 patients sustained a first fragility fracture at some point during the follow-up period. The absolute rate of fracture among gout cases and controls was 53 and 55 per 10 000 person-years, respectively (Table 2). Compared with controls, we found no excess risk of fracture among patients with gout (HR 0.97, 95% CI 0.92–1.02). These findings remained consistent

Table 1: Basic characteristics of the study population

Characteristic	No. (%) of controls* n = 122 961	No. (%) of cases of gout* n = 31 781	Standardized difference†
Age, yr; mean ± SD	63.1 ± 12.2	63.5 ± 12.5	0.04
Male sex	89 978 (73.2)	23 180 (72.9)	0.00
Follow-up time, yr; median (IQR)	10.8 (6.8–13.6)	10.8 (6.7–13.4)	0.03
BMI, kg/m²			
Normal (18.5–24.9)	34 319 (27.9)	5 773 (18.2)	0.41
Underweight (< 18.5)	1 151 (0.9)	129 (0.4)	
Overweight (25.0–29.9)	37 142 (30.2)	11 641 (36.6)	
Obese (≥ 30)	14 852 (12.1)	7 597 (23.9)	
Missing	35 497 (28.9)	6 641 (20.9)	
Smoking status			
Never smoked or former smoker	100 109 (81.4)	26 978 (84.9)	0.09
Current smoker	22 852 (18.6)	4 803 (15.1)	
Alcohol consumption, units per wk			
Never consumed or former drinker	15 143 (12.3)	3 632 (11.4)	0.30
Current 1–9	50 568 (41.1)	12 181 (38.3)	
Current ≥ 10	23 819 (19.4)	9 909 (31.2)	
Unknown	33 431 (27.2)	6 059 (19.1)	
History of falls	3 558 (2.9)	1 143 (3.6)	0.04
Charlson index			
0	96 284 (78.3)	22 419 (70.5)	0.19
1–2	23 377 (19.0)	7 872 (24.8)	
3–4	2 885 (2.3)	1 231 (3.9)	
≥ 5	415 (0.3)	259 (0.8)	
Medications			
Glucocorticoids	3 931 (3.2)	1 506 (4.7)	0.08
Opioids	11 269 (9.2)	4 658 (14.7)	0.17
Bisphosphonates	556 (0.5)	125 (0.4)	0.01
SSRIs	3 979 (3.2)	1 122 (3.5)	0.02
Statins	8 811 (7.2)	3 999 (12.6)	0.18
Antihypertensive agents	29 634 (24.1)	14 233 (44.8)	0.44
Antidiabetic agents	4 894 (4.0)	1 173 (3.7)	0.02
PPIs	8 339 (6.8)	3 366 (10.6)	0.14
NSAIDs	18 892 (15.4)	22 264 (70.1)	1.33
ASA	14 141 (11.5)	5 735 (18)	0.19

Note: ASA = acetylsalicylic acid, BMI = body mass index, IQR = interquartile range, NSAID = nonsteroidal anti-inflammatory drug, PPI = proton pump inhibitor, SD = standard deviation, SSRI = selective serotonin reuptake inhibitor.
*Unless otherwise specified.
†Standardized difference = difference in means or proportion divided by standard error; imbalance defined as absolute value greater than 0.20 (small effect size).

when we stratified our analysis by age. Although women had higher absolute risk of fracture than men, their excess risk compared with their matched controls was not statistically significant (HR 0.96, 95% CI 0.89–1.02). Compared with controls, patients with gout had no increased risk of vertebral or nonvertebral fractures (Supplementary Table 1, Appendix 1). For our 1-year landmark analysis, we included 31 668 patients with incident gout (Figure 2) who did not die, transfer out of the practice or have a fracture within the exposure window. The baseline characteristics of patients who were exposed and those who were not exposed to urate-lowering therapy within 1 year after gout diagnosis are summarized in Supplementary Table 2 (Appendix 1). After propensity score matching, we found no difference in the baseline characteristics by status of exposure to urate-lowering therapy, which shows the quality of our matching (Supplementary Table 3, Appendix 1). There was no difference in the risk of long-term fracture among those who were exposed and were not exposed to at least 6 months of urate-lowering therapy within a year of their diagnosis of gout (HR 1.01, 95% CI 0.84–1.22) (Table 3). We also observed similar findings in our 3-year landmark analysis.

Interpretation

Using data from a large, nationally representative cohort, we compared the risk of fragility fracture (composite of vertebral and nonvertebral fracture) observed among patients with incident gout to the general population. Overall, gout was not associated with an increased risk of vertebral or nonvertebral fracture. These findings were consistent when we stratified our analysis by age and sex. Among those with incident gout, we found that having at least 6 months of urate-lowering therapy within 1 and 3 years of diagnosis had neither adverse nor beneficial effect on the long-term risk of fragility fractures compared with receiving no or less than 6 months of urate-lowering therapy.

We found no association between gout and the risk of fragility fracture. This finding is in contrast to a Taiwanese study that reported a statistically significant 17% higher risk of fracture among patients with gout compared with their matched controls.⁶ This may be a result of the difference in the study population and outcome definition used. For instance, our study focused primarily on fragility fractures, whereas the Taiwanese study included all types of fractures, including fracture of ankle or foot, which accounted for 15% of all fractures in that study's gout cohort and had the largest excess risk (34%). The neutral effect of gout remained when we stratified our analysis by fracture site. This finding is consistent with other studies.^{6,8} Although Tzeng and colleagues⁶ found a 14% increased risk of vertebral fracture, their study failed to take into account important lifestyle-related factors (BMI, smoking status and alcohol consumption), which may have confounded their finding.

Although women are more likely to sustain fragility fractures than men, few studies provide risk estimates by sex.^{6–8} A US-based study²¹ reported positive associations between gout and incidence of hip and wrist fracture in women, with an adjusted excess risk of 12% and 38%, respectively. Although this study refutes our findings, it used self-reported information on both gout and fracture, which may have introduced bias. Furthermore, their study findings may not be generalizable to the wider population, because they were based exclusively on a cohort of nurses.

We found that urate-lowering therapy had neither a beneficial nor adverse effect on the long-term risk of fragility fracture. The existing literature on the topic is conflicting^{6,7,22} and, unlike our study, does not address the issues of immortal time bias or confounding by indication. For instance, Dennison and colleagues⁷ reported 9% excess risk of osteoporotic fractures among those prescribed allopurinol compared with nonusers. Although these authors used propensity score matching, there were still significant differences in the baseline characteristics and comorbidities between groups that were exposed and not exposed. Moreover, their controls may have included

Table 2: Absolute and relative rate of fragility fracture among cases compared with controls*

Characteristic	Unexposed		Exposed		Unadjusted HR (95% CI)	Adjusted HR‡ (95% CI)
	n	Rate† (95% CI)	n	Rate‡ (95% CI)		
Overall	7164	54.7 (53.5–56.0)	1770	52.9 (50.5–55.6)	0.97 (0.92–1.02)	0.95 (0.89–1.01)
Age, yr (quartile)						
1	683	18.3 (16.9–19.7)	179	18.6 (16.1–21.6)	1.02 (0.86–1.20)	0.90 (0.73–1.10)
2	1147	32.2 (30.3–34.0)	296	32.7 (29.2–36.7)	1.01 (0.90–1.16)	0.92 (0.79–1.07)
3	2020	61.7 (59.1–64.5)	505	61.7 (56.6–67.4)	1.02 (0.92–1.12)	1.01 (0.91–1.14)
4	3314	132.4 (127.9–136.9)	790	119.6 (111.6–128.3)	0.92 (0.85–1.00)	0.94 (0.86–1.03)
Sex						
Male	3016	30.9 (29.4–32.0)	793	31.5 (29.4–33.8)	1.02 (0.94–1.10)	0.99 (0.90–1.09)
Female	4148	124.7 (121.0–128.5)	977	117.7 (110.5–125.3)	0.96 (0.89–1.02)	0.94 (0.87–1.01)

Note: BMI = body mass index, CI = confidence interval, HR = hazard ratio.

*Multiple imputation was used to replace missing values of BMI using a chain equation approach based on all baseline characteristics. Five imputed data sets were created and results were combined across all data sets using the Rubin rule to obtain final estimates.

†Per 10 000 person-years.

‡Adjusted for age, alcohol consumption, smoking status, BMI, Charlson index, opioids, fall, glucocorticoids, nonsteroidal anti-inflammatory drugs, acetylsalicylic acid, proton pump inhibitors, antidiabetic and antihypertensive agents, and selective serotonin reuptake inhibitors.

patients who did not have gout, which may have affected their observed association. In contrast, Tzeng and colleagues⁶ reported 28% lower risk of fracture among patients with gout who were prescribed urate-lowering therapy compared with those who were not prescribed the therapy. However, this study overlooked that patients receiving urate-lowering therapy must be event-free from the time of gout diagnosis to the time of the first prescription of urate-lowering therapy be considered exposed, whereas no such requirement is necessary for the unexposed group.

Strengths and limitations

Using data from primary care in the UK, we conducted one of the largest studies with more than 25 years of follow-up to quantify the occurrence of fragility fracture among patients with gout compared with a matched group of individuals without gout in a contemporary population-based manner. Our use of a nationally representative cohort should enable our study findings to be generalizable not only to the UK but also to other developed countries with similar health care systems. Furthermore, the

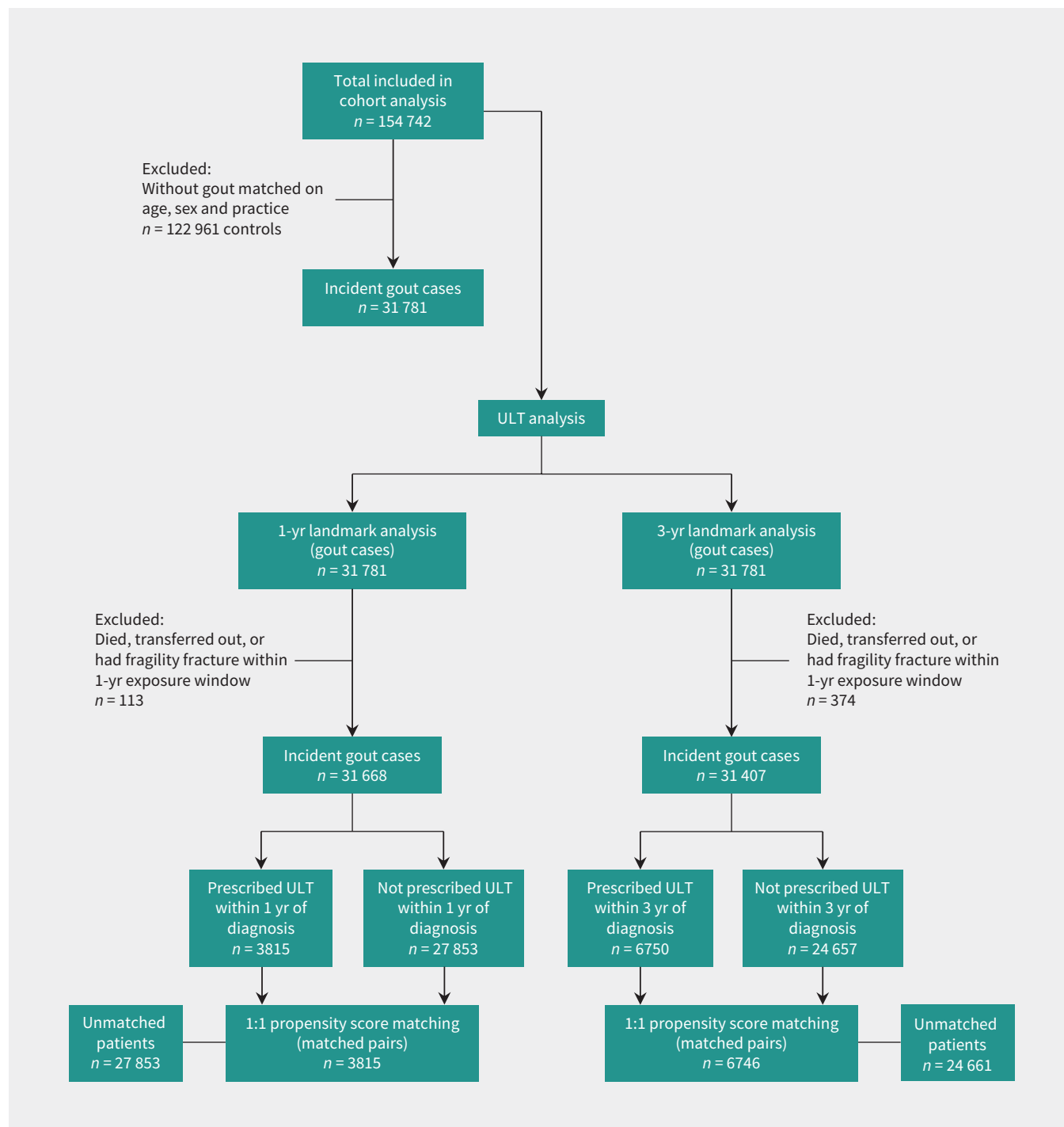


Figure 2: Flow diagram of study population. Note: ULT = urate-lowering therapy.

Table 3: Risk of fracture at 1- and 3-year landmark (propensity score-matched)

Characteristic	1-year landmark analysis			3-year landmark analysis		
	ULT not prescribed rate* (95% CI)	ULT prescribed rate (95% CI)	HR (95% CI)	ULT not prescribed rate (95% CI)	ULT prescribed rate (95% CI)	HR (95% CI)
Overall	62.0 (54.3–70.8)	62.7 (55.0–71.6)	1.01 (0.84–1.22)	65.3 (58.5–72.8)	65.2 (58.5–72.7)	1.00 (0.85–1.16)
Vertebral	8.3 (5.8–12.0)	7.7 (3.7–8.8)	0.68 (0.39–1.20)	9.0 (6.7–12.1)	6.5 (4.6–9.1)	0.71 (0.45–1.12)
Nonvertebral fracture	47.1 (40.4–54.9)	51.9 (44.9–60.0)	1.10 (0.89–1.36)	50.6 (44.7–57.3)	52.1 (46.1–58.9)	1.03 (0.87–1.22)
Wrist	19.2 (15.1–24.4)	21.1 (16.8–26.5)	1.10 (0.79–1.53)	17.9 (14.6–22.1)	18.0 (14.6–22.1)	1.00 (0.74–1.34)
Hip	18.4 (14.4–23.5)	23.1 (18.6–28.7)	1.25 (0.90–1.75)	24.3 (20.3–29.0)	25.2 (21.2–30.1)	1.04 (0.81–1.34)
Humerus	9.5 (6.5–13.3)	7.7 (5.3–11.2)	0.81 (0.49–1.35)	8.4 (6.2–11.4)	8.90 (6.6–11.4)	1.06 (0.70–1.62)

Note: CI = confidence interval, HR = hazard ratio, ULT = urate-lowering therapy.
*Per 10 000 person-years.

prospective nature of the data recording enables us to understand better the temporal relationship between exposure and outcome with minimum bias.

However, our study has several limitations. Our reliance on physicians to make the diagnosis of gout, rather than using the gold standard of visualization of monosodium urate crystals in joint fluid or identification of tophi on examination, could have led to misclassification. However, the diagnosis of gout has been previously validated in the Clinical Practice Research Datalink with high accuracy;¹⁵ it is therefore unlikely that there is any major error in our findings resulting from misclassification of our cases. The findings of this validation study are consistent with another study in which 83% of cases of gout diagnosed by general practitioners were independently validated by a rheumatologist on clinical grounds.²³ Although this previous study found a high positive predictive value, it did not give an indication of negative predictive value (or sensitivity), and there is a possibility of missing cases that may have been diagnosed in a specialized setting. However, this is unlikely, because gout is managed principally in primary care. Furthermore, a recent study using a similar database has shown higher prevalence of gout¹ than previously reported. Therefore, we believe it is unlikely that there is any major error in our findings resulting from misclassification of our cases.

Similar misclassification may also exist for fracture diagnosis, although recording of fracture has also been validated in the Clinical Practice Research Datalink;¹⁶ differential recording in the diagnosis of fracture among cases and controls is unlikely. Finally, the use of 1- and 3-year landmarks for our urate-lowering therapy analysis means that our findings can be generalized only to those who were alive and contributing data at those landmark points, who did not develop fracture within the exposure window and who were prescribed at least 6 months of urate-lowering therapy after their initial diagnosis of gout.

Conclusion

We found no excess risk of fragility fractures among patients with gout. Our findings remained consistent when we stratified our analysis by age, sex and fracture site. Our propensity

score-matched landmark analyses showed that prescription of at least 6 months of urate-lowering therapy within 1 and 3 years of patients' initial diagnosis of gout had neither beneficial nor adverse effects on long-term risk of fracture. These findings should be reassuring to patients, health care policy-makers and clinicians.

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